

Mediation Analysis of Maternal Smoking, Gestational Age, and Birth Weight on the Texas–Mexico Border

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Objectives: Published data on the indirect effect of maternal smoking on birth weight as mediated by gestational age in Hispanic populations are lacking. Our goal was to conduct such a mediation analysis using data from El Paso County, Texas.

Methods: El Paso County is located on the US–Mexico border. A simple mediation analysis was conducted using year 2010 El Paso County birth certificate data. The SAS macro PROCESS 3.5.3 was used to estimate the direct and indirect effects of active maternal smoking (by trimester) on birth weight (in grams) in the setting of linear regression. The single mediator was gestational age in weeks. A direct or indirect effect was deemed to be present if the 95% confidence limits (CLs) excluded 0. Analyses were adjusted for multiple variables, including maternal prepregnancy body mass index. The indirect effect was reported along with a 95% bootstrap CL.

Results: A total of 16,654 singleton births were included in the cohort. The majority of the mothers were White Hispanic (87.2%). The mean (standard deviation) birth weight was 3198.6 g (517.2). A direct effect of maternal smoking during each trimester on birth weight was detected. An indirect effect of maternal smoking on birth weight was not detected in any of the trimesters. In adjusted analyses for the third trimester, the indirect effect for every 1-U increase in the mean number of cigarettes smoked per day was –4.18 (95% bootstrap CL –10.64 to 1.99).

Conclusions: In our large, predominantly Hispanic cohort, it appears that gestational age is not a mediator of the effect of maternal smoking on birth weight. Future studies in our population should explore other possible mediators of this association.

Key Words: birth weight, gestational age, Hispanic ethnicity, maternal smoking, mediation analysis

Preterm birth (birth before 37 completed weeks of gestation) and low birth weight are conditions of major public health importance. According to the March of Dimes, 10.2% of the infants (live births) born in the United States in 2019 were preterm.¹ In 2019 the incidence of low birth weight (defined as a birth weight < 2500 g) was 8.3% of live births.²

The prevalence of maternal smoking during pregnancy in the United States was 9.2% in 2010, and this measure decreased steadily to 6.9% in 2017.³ Maternal cigarette smoking during pregnancy has been linked to an increased risk of preterm birth and low birth weight.⁴ A recent meta-analysis of studies conducted in the Americas found an odds ratio of 2.00 (95% confidence interval 1.77–2.26) for the association between active maternal smoking and low birth weight.⁵

Certain previous investigations of non-Hispanic White and Black women (summarized by Singleton et al⁶) have noted that the association between maternal smoking and low birth weight varied by race; however, similar data are lacking in Hispanic populations. A PubMed search did not reveal any publications reporting both the direct effect of maternal smoking during pregnancy on birth weight and the indirect effect of maternal smoking on birth weight mediated by gestational age in a predominantly Hispanic population. This gap in knowledge was

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Key Points

- Active maternal smoking during pregnancy is associated with an increased incidence of low birth weight, and this relation has been reported to vary by race (non-Hispanic White and Black).
- Statistical mediation analyses in Hispanic populations in which the risk factor is maternal smoking, the mediator is gestational age, and the outcome is birth weight are needed.
- In our large, predominantly Hispanic sample, gestational age was not a mediator of the effect of maternal smoking on birth weight.

addressed using birth certificate data from El Paso County, Texas. El Paso County is located on the Texas–Mexico border, and its population was estimated to be 839,238 in 2019.⁷ The majority of the residents (82.9%) of El Paso County identify as Hispanic or Latinx.⁷

Methods

The study protocol was reviewed by the Institutional Review Board for the Protection of Human Subjects at Texas Tech University Health Sciences Center El Paso and was deemed exempt from formal institutional review board review.

Source of Data and Inclusion Criteria

A retrospective cohort study was conducted using deidentified birth certificate data from El Paso County, Texas, for the year 2010. The overall dataset has information on 17,503 births. Our analysis was restricted to singleton births. Records with an unknown value for gestational age, any of the maternal cigarette smoking variables, or maternal level of education were deleted. If a record had a gestational age of >44 weeks (an implausible value), then it was excluded from our analysis.

Statistical Analysis

Data on key variables were summarized using means and standard deviations (SDs) for continuous variables and number and percentage for categorical and count variables. Preterm birth was defined as birth before 37 completed weeks of gestation. Low birth weight was defined as an infant birth weight < 2500 g. A mediation analysis was performed using SAS 9.4 software (SAS Institute, Cary, NC) and the SAS macro PROCESS version 3.5.3 (authored by Andrew F. Hayes).⁸ The Figure displays a causal diagram known as a directed acyclic graph, which illustrates the hypothesized relations between the variables that were studied.⁹ Maternal smoking during pregnancy was the independent variable (X) and the outcome (Y) was birth weight in grams. Gestational age served as the single mediator of interest and is represented by the letter M . The total effect, direct effect, and indirect effect of maternal smoking on birth weight were quantified using a simple mediation model in the setting of linear regression. Maternal smoking was recorded in the dataset as the average number of cigarettes smoked per day.

Analyses were stratified by maternal smoking during each trimester.

Let a , b , c , and c' represent causal effects (Fig.; note that the quantity c is not shown in the Figure). These four quantities were parameter estimates (regression coefficients) from linear regression models. The total effect of X on Y is denoted by c and is defined as the direct effect (c') plus the indirect effect. The indirect effect is the product of a and b :

$$c = c' + (a)(b) \quad (1)$$

The four quantities in Equation 1 were derived from the following three ordinary least squares linear regression models

$$\hat{Y} = \beta_0 + cX, \quad (2)$$

$$\hat{M} = \beta_0 + aX, \quad (3)$$

and

$$\hat{Y} = \beta_0 + c'X + bM. \quad (4)$$

The direct effect of maternal smoking during pregnancy is the unmediated effect of this exposure on birth weight. The indirect effect of maternal smoking during pregnancy is the portion of the exposure that is mediated through gestational age.

Given that the sampling distribution of the indirect effect is typically not Gaussian distributed, calculating confidence limits (CLs) for the true indirect effect based on normality is not appropriate.⁸ As such, statistical inference about the indirect effect of maternal smoking was performed by inspecting the 95% bootstrap CL produced by the PROCESS macro using the percentile method.⁸ A total of 12,000 bootstrap samples were requested from PROCESS. If the 95% bootstrap CL excludes the value of 0, then we can conclude with 95% confidence that an indirect effect different from 0 is present. Parameter estimates were adjusted for maternal age, number of prenatal care visits, maternal race and ethnicity, educational level, nulliparity, diabetes mellitus (prepregnancy or gestational), and prepregnancy maternal body mass index.

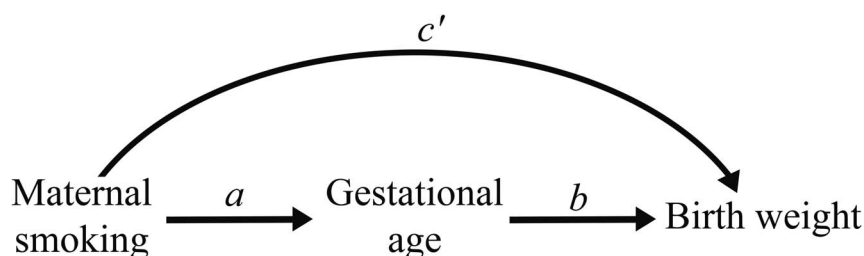


Fig. Directed acyclic graph for the hypothesized effect of maternal smoking during pregnancy on birth weight, where a , b , and c' represent causal effects estimated from linear regression models.

Results

A total of 16,654 records were available for analysis after applying our inclusion and exclusion criteria. Characteristics of the cohort are reported in Table 1. The mean maternal age (SD) was 26.3 years (6.2 years) and the majority of the mothers were White Hispanic (87.2%). The prevalence of smoking on average one to nine cigarettes per day during the first trimester was 1.6% (Table 1). This prevalence decreased to 1.0% in the third trimester. Gestational age ranged from 18 to 44 completed weeks, and the incidence of preterm birth (before 37 completed weeks) was 12.3% (n = 2043). The mean birth weight (SD) was 3198.6 g (517.2 g) (Table 1). The incidence of low birth weight (defined as a birth weight < 2500 g) was 7.3% (n = 1208).

Table 2 presents the adjusted results of the simple mediation analysis. The linear regression parameter estimates were adjusted for maternal age, the number of prenatal care visits, maternal race-ethnicity, maternal educational level, nulliparity, diabetes mellitus (prepregnancy or gestational), and maternal body mass index. A direct effect of maternal smoking during each trimester on birth weight was detected. In the third trimester, for every 1-U increase in the mean number of cigarettes smoked per day, there was a decrease in birth weight of 14.82 g (95% CL –25.02 to –4.62). An indirect effect of maternal smoking was not detected in any of the trimesters. In the third trimester the indirect effect of the mean number of cigarettes smoked per day was –4.18 (95% bootstrap CL –10.64 to 1.99).

Discussion

In a simple mediation analysis, the indirect effect estimates the impact of *X* on *Y* through a single intermediary.⁸ Our simple mediation analysis in a predominantly Hispanic population on the US–Mexico border did not detect an indirect effect of maternal smoking on infant birth weight via gestational age.

The strengths of our study include its large sample size and uniqueness. We could not find a similar investigation that had been conducted in El Paso County, Texas. An additional strength of our investigation was the use of modern mediation analysis techniques rather than the causal steps approach that was popularized in the 1980s by Baron and Kenny. It has been noted that the Baron and Kenny method has several limitations and flaws.⁸

A limitation of this study is the unknown accuracy (eg, sensitivity, specificity, positive predictive value) of the variables found in this county birth certificate dataset. We could not find any published studies that reported measures of accuracy for El Paso County, Texas, birth certificate data. A study by Northam et al investigated the birth certificate data collection methods at five high-delivery hospitals in North Texas.¹⁰ The authors found “considerable variance in how birth certificate data were obtained” at these five hospitals.¹⁰ Of note, these hospitals were visited in 2001, before the adoption of the 2003 revised US standard birth certificate. Our study focused on year 2010 data.

Table 1. Characteristics of 16,654 singleton births in El Paso County, Texas, 2010

Characteristic	No. (%) or mean (SD)
Demographic variables	
Maternal age, y, mean (SD)	26.3 (6.2)
Maternal race and ethnicity, N (%)	
White Hispanic	14,514 (87.2)
White non-Hispanic	1435 (8.6)
Other	705 (4.2)
Maternal educational level, N (%)	
Bachelor’s degree or higher	3107 (18.7)
Clinical and obstetric variables	
Body mass index (kg/m ²), mean (SD)	25.9 (5.7)
Nulliparous, N (%)	6566 (39.4)
No. prenatal visits, N (%)	
0	489 (2.9)
1–4	1227 (7.4)
5–8	5154 (31.0)
≥9	9784 (58.8)
Diabetes mellitus, N (%)	
Pregpregnancy diabetes mellitus	45 (0.3)
Gestational diabetes mellitus	345 (2.1)
Pregpregnancy or gestational diabetes mellitus	390 (2.3)
Main exposure variables	
Mean no. cigarettes smoked per day by mother during first trimester, N (%)	
0	16,312 (98.0)
1–9	268 (1.6)
10–19	58 (0.4)
≥20	16 (0.1)
Mean no. cigarettes smoked per day by mother during second trimester, N (%)	
0	16,381 (98.4)
1–9	216 (1.3)
10–19	47 (0.3)
≥20	10 (0.1)
Mean no. cigarettes smoked per day by mother during third trimester, N (%)	
0	16,459 (98.8)
1–9	158 (1.0)
10–19	32 (0.2)
≥20	5 (0.03)
Mediator	
Gestational age, wk, mean (SD)	38.4 (2.2)
Outcome	
Birth weight, g, mean (SD)	3198.6 (517.2)

SD, standard deviation.

The Texas Department of State Health Services’ Center for Health Statistics (CHS) was contacted by the corresponding author and was informed that the CHS has not conducted validation

Table 2. Adjusted^a results of the simple mediation analyses derived from linear regression models stratified by trimester

Trimester	Direct effect of MS		Indirect effect of MS	
	Estimate	95% CL	Estimate	95% CL
First	-13.18	-19.87 to -6.50	-1.91	-6.74 to 2.81
Second	-12.80	-19.90 to -5.70	-1.19	-8.35 to 4.04
Third	-14.82	-25.02 to -4.62	-4.18	-10.64 to 1.99

The exposure variable is MS reported as the mean number of cigarettes smoked per day in each trimester, the mediator is gestational age in weeks, and the outcome is infant birth weight. CL, confidence limit; MS, maternal smoking.

^aParameter estimates were adjusted for maternal age, number of prenatal care visits, maternal race and ethnicity, educational level, nulliparity, diabetes mellitus (prepregnancy or gestational), and maternal body mass index.

studies of the accuracy of Texas birth certificate data (CHS Vital Events Data Management, personal communication). All of the counties in Texas follow state guidelines in regard to registering births (Field Services representative, Texas Department of State Health Services, Vital Statistics Section, personal communication), however. Ninety-nine percent of all births in Texas occur in a hospital, and birth registrars are required to complete the birth certificate within 5 days (Field Services representative, Texas Department of State Health Services, Vital Statistics Section, personal communication). In addition, the Texas Department of State Health Services provides online resources to assist birth registrars in collecting accurate data,¹¹ and this agency hosts trainings for birth registrars during their conferences (CHS Vital Events Data Management, personal communication).

A discussion of maternal smoking is merited. The prevalence of active maternal smoking during pregnancy in our cohort was lower than expected. Given the social stigma of maternal smoking, smoking status may have been underreported in our sample.¹² Ideally, active smoking by the mother during pregnancy would be validated using cotinine values.¹³ Cotinine can be measured in a variety of specimens, including saliva, serum, and urine.¹⁴ A major limitation of our study is that information on this gold standard of smoking status was not found in the birth certificate dataset.

Detailed information on preterm birth also was lacking in our database. For example, it is unknown whether a preterm birth was iatrogenic or spontaneous. Given this limitation and those listed earlier, our findings have limited clinical significance. Nonetheless, the authors hope to stimulate a dialog on the role of mediation analysis in medicine. Clinicians and health

scientists interested in applying mediation analysis to real-world data are referred to the article by Jung.¹⁵

Conclusions

This mediation analysis using birth certificate data did not find an effect of maternal smoking on infant birth weight through our hypothesized mediator of gestational age. Future similar studies should consider other possible mediators, such as intrauterine growth restriction, and also strive to collect information on maternal passive smoking.

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